

# HOW RANDOM ARE THESE GRASSES?

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# A FEW STEPS BEYOND RANDOM

**Main claim:** Many cases of both ecological and mathematical interest have *very*, but *not completely*, random(-like) interactions.

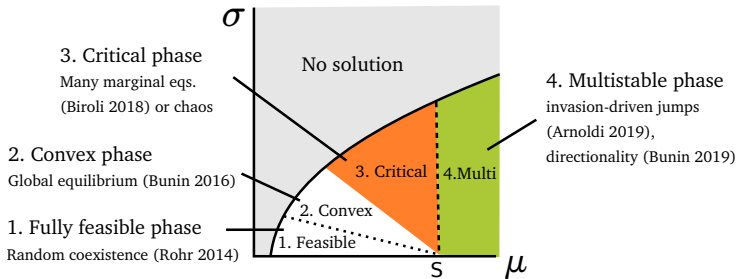
What are the interesting first steps beyond fully-random matrices?

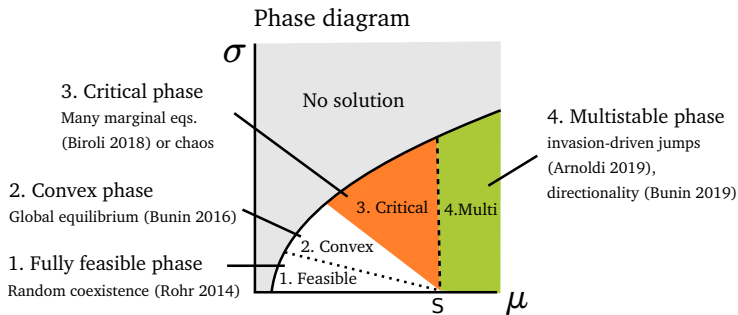
## Plan:

- 1 Minimal structure allowing coexistence
- 2 Dynamically-relevant extra structure (if time allows)

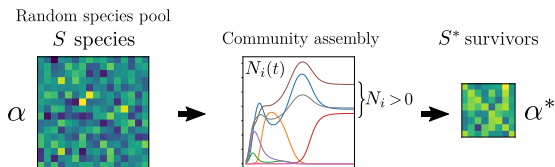
## Part 1: Minimal structure

# Phase diagram





In region 2, LV dynamics go to globally stable equilibrium, but only  $S^* < S$  species survive in it.



Feasible submatrix  $\alpha^*$  elements are **not i.i.d.** but they have the *minimal deviation from randomness* that ensures feasibility

## Subplan:

- Explain the correlation structure
- Show it in data from grassland experiments

- Our notations for competitive Lotka-Volterra ( $\alpha > 0 =$  negative interaction)

$$\frac{dN_i}{dt} = \frac{r_i}{K_i} N_i \left( K_i - N_i - \sum_{j \neq i}^S \alpha_{ij} N_j \right) \quad (1)$$

- Equilibrium condition for the  $S^*$  survivors

$$N_i = K_i - \sum_{j \neq i}^{S^*} \alpha_{ij}^* N_j \quad (2)$$

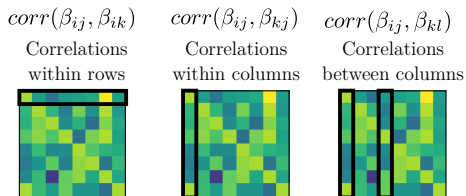
- For empirical reasons (explained later), we rather use rescaled abundances  $\eta_i = N_i/K_i$

$$\eta_i = 1 - \sum_{j \neq i}^{S^*} \beta_{ij}^* \eta_j \quad (3)$$

NB:  $\beta_{ij}^* = \alpha_{ij}^* K_j / K_i$

# CORRELATIONS

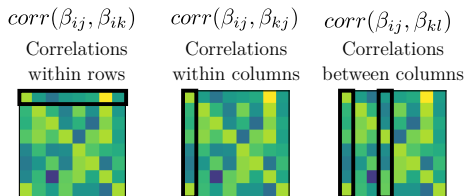
Bunin (PRE 2017) showed that  $\beta_{ij}^*$  are not i.i.d.  
Species die = rows and columns disappear, creating biases and correlations  
(shown later).





# CORRELATIONS

Bunin (PRE 2017) showed that  $\beta_{ij}^*$  are not i.i.d.  
Species die = rows and columns disappear, creating biases and correlations  
(shown later).



Surprisingly, same result from simple probabilistic argument:  
rather than run dynamics, directly draw  $\beta_{ij}$  from Gaussian with mean  $\bar{\beta}$  subject to  
the  $S$  linear constraints

$$0 = 1 - \eta_i - \sum_{j \neq i}^S \beta_{ij} \eta_j \quad (4)$$

# CORRELATIONS

Simple calculation (Lagrange multipliers or rotation) gives:

$$E[\beta_{ij}|\eta_i, \eta_j] = \bar{\beta} + (1 - \bar{\beta})\Delta(\eta_i, \eta_j) \quad (5)$$

$$\text{corr}(\beta_{ij}, \beta_{ik}|\eta_i, \eta_j, \eta_k) = -\frac{\eta_j\eta_k}{\sum_{m \neq i} \eta_m^2}. \quad (6)$$

where mean interactions are biased to achieve the correct  $\eta_i$ :

$$\Delta(\eta_i, \eta_j) = -\frac{(\eta_i - \eta^*)\eta_j}{\sum_{m \neq i} \eta_m^2}. \quad (7)$$

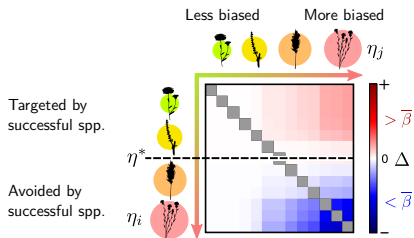
and  $\eta^*$  is baseline abundance obtained if all  $\beta_{ij} = \bar{\beta}$

$$\eta^* = \frac{1 - \bar{\beta} \sum_i \eta_i}{1 - \bar{\beta}}. \quad (8)$$

# CORRELATIONS

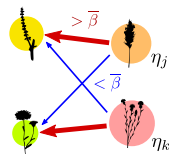
## Two statistical patterns among coexisting species

(a) Competitive bias  $\Delta(\eta_i, \eta_j) \sim -(\eta_i - \eta^*)\eta_j$

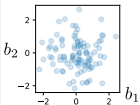
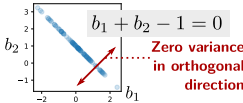
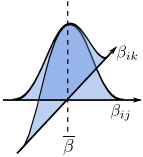
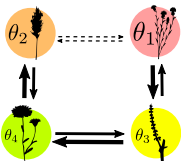
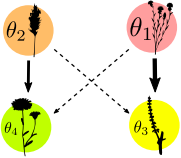
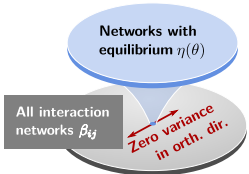


(b) Overlap avoidance  $\text{CORR}(\beta_{ij}, \beta_{ik}) \sim -\eta_j\eta_k$

Two successful species do not target the same species.

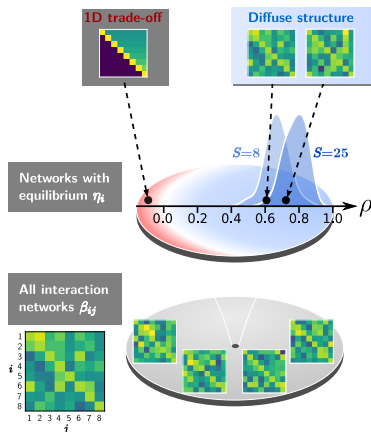


# HOW DOES IT WORK?

	Prior distribution	Bias on means	Correlation/covariance	Transformed distribution
Simple example		$\text{mean}(b_1)=0$ $\text{mean}(b_2)=1$	$\text{corr}(b_1, b_2)=-1$ $\text{Cov}_{ij} = \begin{pmatrix} 1 & -1 \\ -1 & 1 \end{pmatrix}$	 <p><math>b_1 + b_2 - 1 = 0</math></p> <p>Zero variance in orthogonal direction</p>
Clique pattern	<p>Multivariate Gaussian</p>  <p>Prior mean <math>\bar{\beta}</math></p>	$\text{mean}(\beta_{ij}) \sim -A\theta_i\theta_j$ 	$\text{corr}(\beta_{ij}, \beta_{ik}) \sim -A\theta_j\theta_k$ 	$\sum_j \beta_{ij}\theta_j - 1 = 0$  <p>Networks with equilibrium <math>\eta(\theta)</math></p> <p>All interaction networks <math>\beta_{ij}</math></p> <p>Zero variance in orth. dir.</p>

# INTERPRETATION

Many interaction matrices  $\beta_{ij}$  admit the same equilibrium  $\eta_i$   
( $O(S^2)$  parameters for only  $S$  constraints)



Most (from random prior) will exhibit our predicted correlations,  
i.e. minimal deviation from randomness allowing these  $\eta_i$

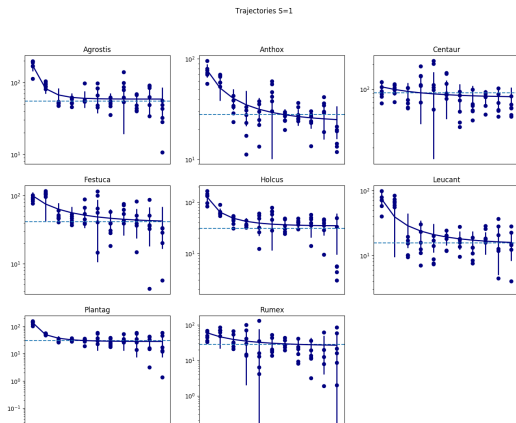
# FINDING IT IN DATA



Each plot = some combination of 1, 2... species from a pool of  $S$  species

# WHAT DATA?

Data = time series  $N_i(t)$  in various combinations of species



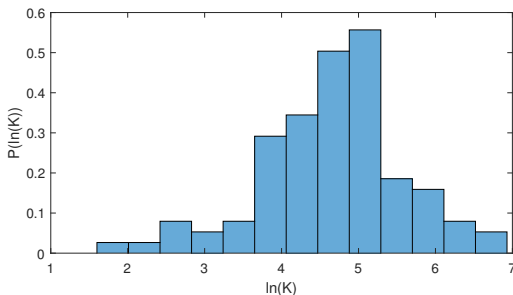
(here Wageningen grassland experiment,  $S = 8$ , 11 years, 4 replicates)

# INFERRING INTERACTIONS

- Somehow estimate equilibrium abundance  $N_i$
- Multilinear fit of interactions  $\alpha$ :  
each species combination  $k$  = point on hyperplane defined by

$$N_i^{(k)} = K_i - \sum_{j \in k} \alpha_{ij} N_j^{(k)} \quad (9)$$

- Problem:  $K_i$  are widely distributed (and  $\alpha$  too)



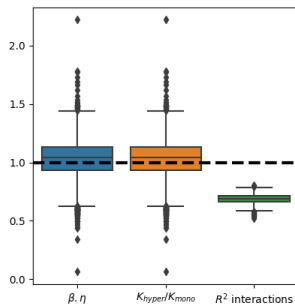


# INFERRING INTERACTIONS

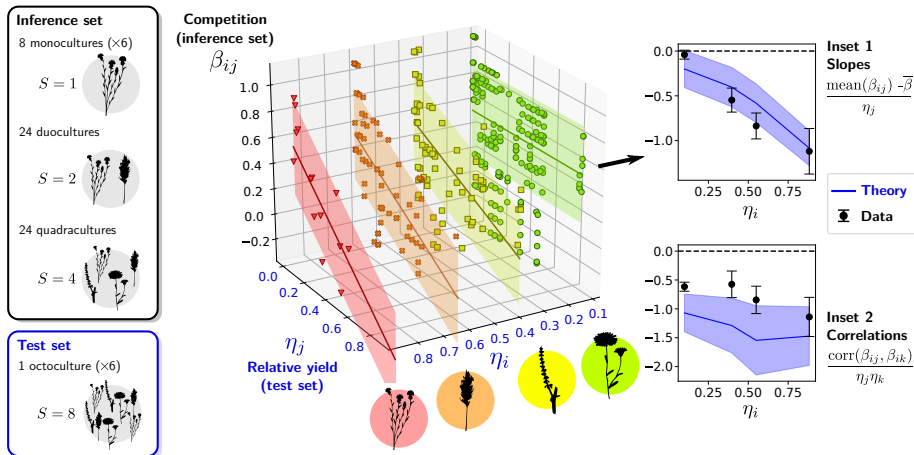
- Solution: take  $\eta_i = N_i/K_i$  and fit

$$\eta_i = 1 - \sum_{j \neq i} \beta_{ij} \eta_j \quad (10)$$

- Sanity checks: Wageningen really looks like LV equilibrium



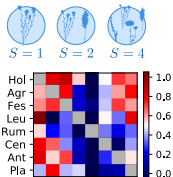
Once we have individual interactions  $\beta_{ij}$ , see if they show the expected trends



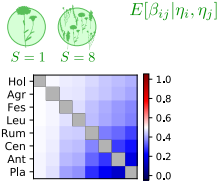
(NB: theoretical predictions controlled only by  $\eta_i$  in the full  $S = 8$  composition; shaded area = variation of prediction due to error on  $\eta_i$ )

Every way we slice and dice the data, it matches theory

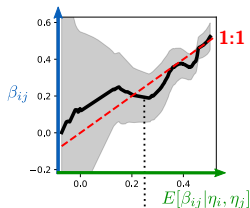
**(a) Interaction coefficients  $\beta_{ij}$**



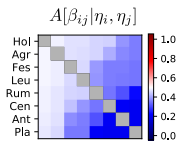
**(b) Theoretical expectations**



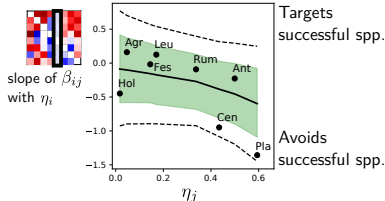
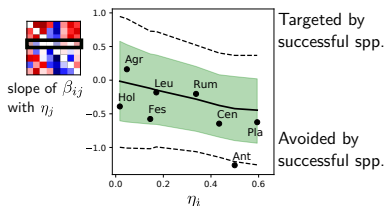
**(c) One-to-one comparison**



**(d) Empirical average**



**(e) Per-species comparison**

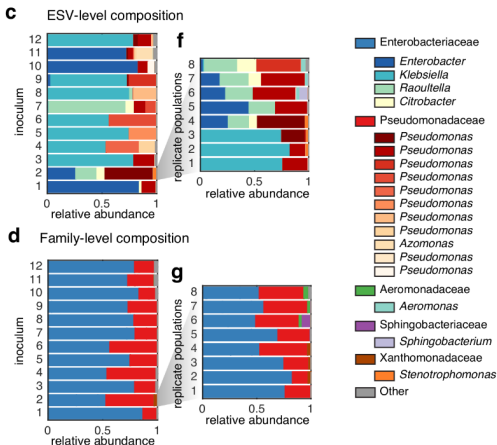


Rightmost panel: shaded area:  $\pm 1$  SD of predictions, dashes: 95% CI of predictions

- Within a group of sufficiently similar grasses, good empirical evidence of minimal deviation from randomness (just enough to allow coexistence)
- What about larger ecological networks, how random are they?

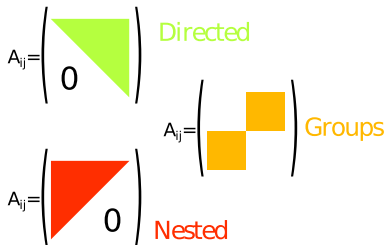
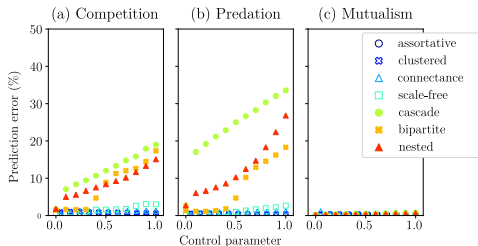
## Part 2: Extra structure

- Random matrix: all species are statistically equivalent, hence we care only about scaled mean  $\mu$ , SD  $\sigma$  and symmetry  $\gamma$  of interactions.
- Reality: not all statistically equivalent, but there are structures e.g. functional groups



(Goldford et al)

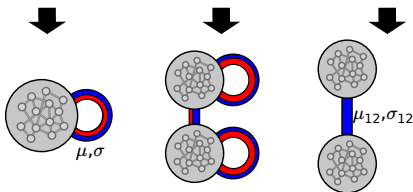
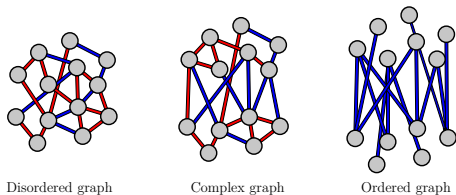
- Not all extra structure is dynamically relevant
- In ecological models, large macrostructures are main causes of departure from mean-field (random prediction)



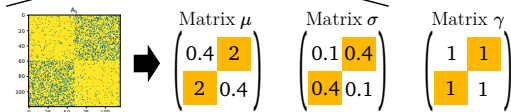
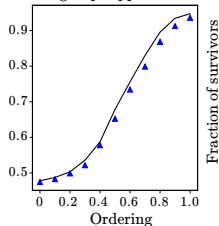
x-axis = degree of order (structure-specific control parameter)  
 y-axis = difference between abundance distribution for random versus structured interactions

(Barbier, Arndli, Bunin and Loreau, PNAS 2018)

Simple extra structure can be captured by a not-completely-random approximation, e.g. matrix of per-group interaction mean  $\mu_{xy}$  and SD  $\sigma_{xy}$

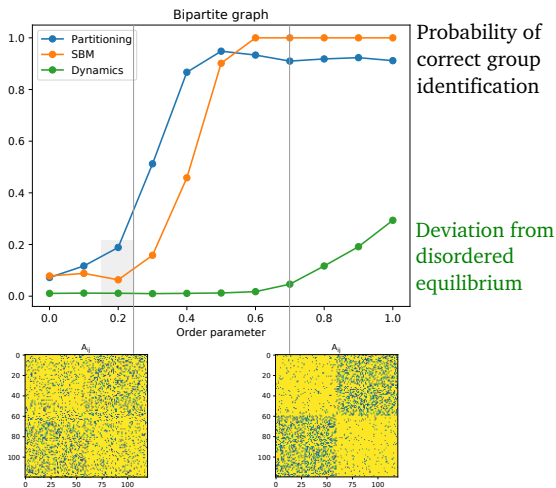


Two groups approximation





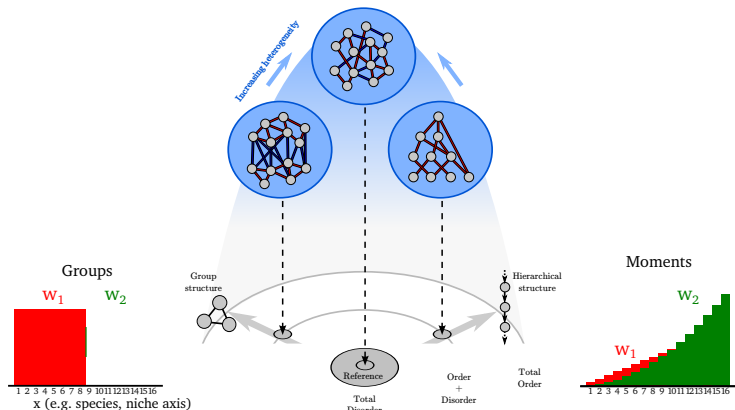
# Stochastic Block Models identify the correct structure in the interaction network



(in fact they are a bit overzealous: they identify it *before* it starts to matter in susceptibility  $V_{ij}$ , which is more self-averaging)

# DIRECTIONS OF DEVIATION FROM RANDOMNESS?

Different ways of adding order with more parameters: groups (matrix  $\mu_{xy}$ ), Taylor expansion (function  $\mu(x, y)$ )...



Technical point of view: low rank structures, perturbed random matrix theory

# CONCLUSIONS

- Random matrices are not a bad starting point for real systems: empirical evidence of minimal deviation from randomness
- However it is worth looking a bit beyond randomness, into simple order+disorder combinations

*"There is a fundamental dichotomy between structure and randomness, which in turn leads to a decomposition of any object into a structured (low-complexity) component and a random (discorrelated) component."*

Terence Tao, The dichotomy between structure and randomness, arithmetic progressions, and the primes. 2006 ICM proceedings.