

Tutorial on Random Matrix Theory

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Quick reminder of the motivations

The paradigm of May (~ 1971)

General idea : for complex multispecies networks, detailed interactions are rarely available, random matrix models may be relevant.

More precisely, an ecological network is modelled as a system of ODEs of the form:

$$\frac{da_i(t)}{dt} = f_i(a_1, \dots, a_N), \quad 1 \leq i \leq N,$$

with a_i the abundance of species i .

Assume the existence of an equilibrium (a_1^*, \dots, a_N^*) and linearise near equilibrium. Namely, if we set $a_i(t) = a_i^* + x_i(t)$, the dynamics can be approximated at first order by the linear system:

$$\frac{d}{dt} \mathbf{x}(t) = \mathcal{J}(\mathbf{a}^*) \cdot \mathbf{x}(t),$$

with $\mathcal{J}(\mathbf{a}^*)_{k\ell} := \frac{\partial f_k}{\partial x_\ell}(a_1^*, \dots, a_N^*)$.

May's model : the entries of \mathcal{J} are i.i.d. random variables (except on the diagonal).

Lotka-Volterra models

The dynamics of interacting species may be described by the Lotka-Volterra equations:

$$\frac{da_i(t)}{dt} = a_i \left(r_i - \theta a_i + \sum_{j=1}^N \frac{Z_{ij}}{N^\delta} a_j \right)$$

where

- ▶ r_i is the intrinsic growth rate of species i
- ▶ $\theta > 0$ is a friction coefficient (intraspecific competition)
- ▶ Z_{ij} stands for the interactions $j \rightarrow i$.

The equilibrium \mathbf{a}^* is given by

$$\mathbf{a}^* = \left(\theta I_N - \frac{Z}{N^\delta} \right)^{-1} \mathbf{r}$$

The Jacobian $\mathcal{J}(\mathbf{a}^*)$, which is explicit

$$\mathcal{J}(\mathbf{a}^*) = \text{diag}(\mathbf{a}^*) \left(-\theta I_N + \frac{Z}{N^\delta} \right)$$

It is a special case of *density-dependent models*, for which

$$\mathcal{J}(\mathbf{a}^*) = \text{diag}(\mathbf{a}^*) \cdot \mathbf{M},$$

where \mathbf{M} may depend on $\text{diag}(\mathbf{a}^*)$.

Large Random Matrices

Random matrices

It is a $N \times N$ matrix

$$\mathbf{Y}_N = \begin{bmatrix} Y_{11} & \cdots & Y_{1N} \\ \vdots & & \vdots \\ Y_{N1} & \cdots & Y_{NN} \end{bmatrix}$$

whose entries ($Y_{ij}; 1 \leq i, j \leq N$) are random variables.

Matrix features

Of interest are the following quantities

- ▶ **\mathbf{Y}_N 's spectrum** ($\lambda_i, 1 \leq i \leq N$) and eigenvectors
- ▶ **Extreme eigenvalues** (λ_{\min} and λ_{\max} if spectrum is real, spectral radius, eigenvalue with maximal real part etc.)
- ▶ **Linear statistics** $\text{Trace } f(\mathbf{Y}_N) = \sum_{i=1}^N f(\lambda_i)$ and their fluctuations

Asymptotic regime

Often, the description of the previous features **takes a simplified form as**

$$\boxed{N \rightarrow \infty}$$

Wigner Matrices

Matrix model

Let $\mathbf{X}_N = (X_{ij})$ a real symmetric (or Hermitian) $N \times N$ matrix with i.i.d. entries **on and above** the diagonal with

$$\mathbb{E}X_{ij} = 0 \text{ and } \mathbb{E}|X_{ij}|^2 = 1$$

and $X_{ij} = \overline{X_{ji}}$ (for symmetry).

- ▶ consider the spectrum of **Wigner**

matrix $\mathbf{Y}_N = \frac{\mathbf{X}_N}{\sqrt{N}}$

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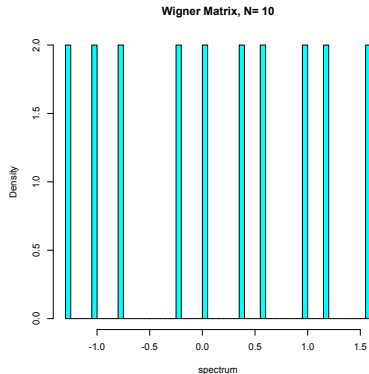


Figure: Histogram of the eigenvalues of \mathbf{Y}_N

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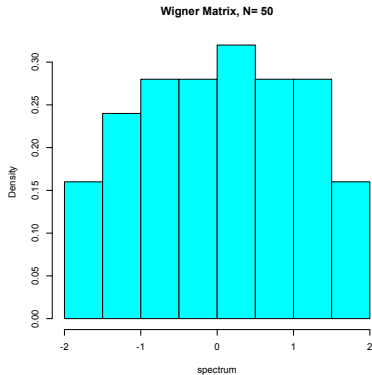


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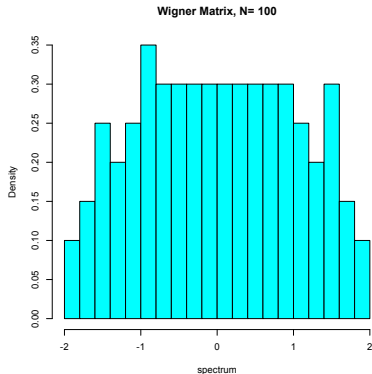


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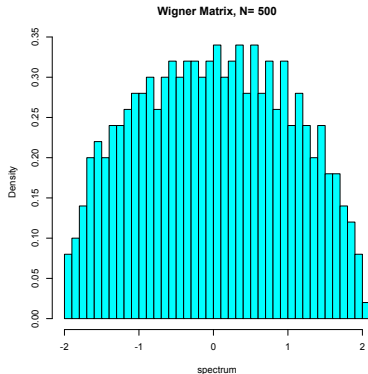


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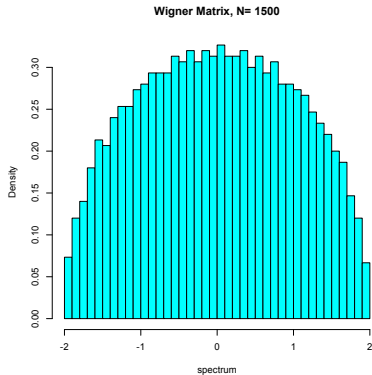


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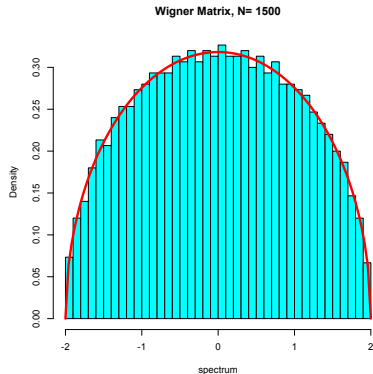


Figure: The semi-circular distribution (in red) with density $x \mapsto \frac{\sqrt{4-x^2}}{2\pi}$

Wigner's theorem (1948)

"The histogram of a Wigner matrix converges to the **semi-circular distribution**"

About normalisation

Wigner's theorem can be written as follows :

$$\frac{1}{N} \sum_{i=1}^N f(\lambda_i(\mathbf{Y}_N)) \xrightarrow[N \rightarrow \infty]{a.s.} \int_{-2}^2 f(x) \frac{\sqrt{4-x^2}}{2\pi} dx = \int f(x) d\sigma(x).$$

A quick computation

$$\begin{aligned} \frac{1}{N} \sum_{i=1}^N \lambda_i(\mathbf{Y}_N)^2 &= \frac{1}{N} \text{Trace}(\mathbf{Y}_N)^2 = \frac{1}{N^2} \text{Trace}(\mathbf{X}_N)^2 \\ &= \frac{1}{N^2} \sum_{i,j=1}^N X_{ij} X_{ji} = \frac{1}{N^2} \sum_{i,j=1}^N |X_{ij}|^2 \end{aligned}$$

so that

$$\mathbb{E} \left(\frac{1}{N} \sum_{i=1}^N \lambda_i(\mathbf{Y}_N)^2 \right) = 1.$$

In the Lotka-Volterra model, this normalisation corresponds to $\delta = 1/2$.

Fluctuations of linear statistics

Many results starting with Johansson (1999) lead to the following results: if the entries have enough finite moments,

- ▶ for smooth test functions f ,

$$N \left(\frac{1}{N} \sum_{i=1}^N f(\lambda_i(\mathbf{Y}_N)) - \int f(x) d\sigma(x) \right)$$

converges to a Gaussian variable, with mean and variance depending on f .

- ▶ for $f(x) = \mathbf{1}_{[y, \infty)}(x)$, we let $N_y(\mathbf{Y}_N) := \#\{i; \lambda_i(\mathbf{Y}_N) \geq y\}$. If the first fourth moment match with a normal variable, $\text{Var}(N_y(\mathbf{Y}_N)) = \left(\frac{1}{2\pi^2} + o(1)\right) \log N$, and

$$\frac{N_y(\mathbf{Y}_N) - N\sigma([y, \infty))}{\sqrt{\text{Var}(N_y(\mathbf{Y}_N))}}$$

converges to a normal random variable.

Local spectrum behavior for Wigner matrices

Let

$$\lambda_{\max} = \lambda_{\max}(\mathbf{Y}_N) \quad \text{and} \quad \lambda_{\min} = \lambda_{\min}(\mathbf{Y}_N)$$

Theorem (convergence of extremal eigenvalues)

If $\mathbb{E}|X_{ij}|^4 < \infty$, then

$$\lambda_{\max} \xrightarrow[N \rightarrow \infty]{\text{a.s.}} 2 \quad \text{and} \quad \lambda_{\min} \xrightarrow[N \rightarrow \infty]{\text{a.s.}} -2$$

Theorem (fluctuations of λ_{\max} and Tracy-Widom distribution)

If $\lim_{s \rightarrow \infty} s^4 \mathbb{P}(|X_{12}| \geq s) = 0$, we have

$$N^{2/3} \{\lambda_{\max}(\mathbf{Y}_N) - 2\} \xrightarrow[N \rightarrow \infty]{\mathcal{L}} \mathbb{P}_{\text{TW}}.$$

Remark

- ▶ Notice the **strange normalization** $N^{2/3}$
- ▶ This stems from the **\sqrt{x} -shape of the density at the edge** (see later)

Details on Tracy-Widom distribution

Tracy-Widom distribution is defined by

- ▶ its cumulative distribution function

$$F_{TW}(x) = \exp \left\{ - \int_x^\infty (u-x)^2 q^2(u) du \right\}$$

- ▶ where

$$q''(x) = xq(x) + 2q^3(x) \quad \text{and} \quad q(x) \sim \text{Ai}(x) \text{ as } x \rightarrow \infty .$$

$x \mapsto \text{Ai}(x)$ being the Airy function.

Don't bother .. just download it

- ▶ For simulations, cf. R Package 'RMTstat', by Johnstone et al.
- ▶ Also, Folkmar Bornemann (TU München) has developed fast `matlab` code

Tracy-Widom curve

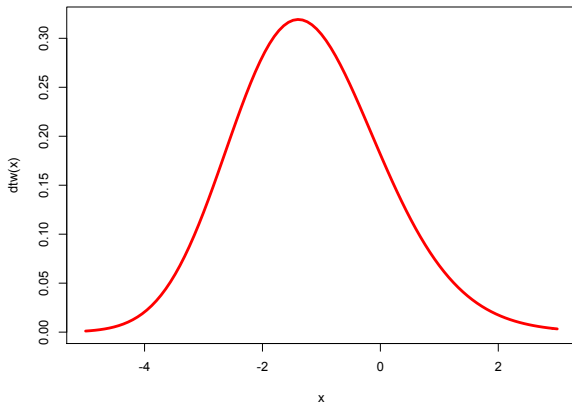


Figure: Tracy-Widom density

A heuristic on the normalization $N^{2/3}$

► By Wigner's theorem, $\frac{\#\{\lambda_i > 2 - \varepsilon\}}{N} \rightarrow \int_{2-\varepsilon}^2 \frac{\sqrt{4-x^2}}{2\pi} dx$

► Hence "for small ε ",

$$\begin{aligned}\#\{\lambda_i > 2 - \varepsilon\} &\approx N \int_{2-\varepsilon}^2 \frac{\sqrt{4-x^2}}{2\pi} dx \\ &\approx N \frac{4}{2\pi} \int_{2-\varepsilon}^2 \sqrt{2-x} dx = cN\varepsilon^{3/2}\end{aligned}$$

► To have **finitely many values** in $(2 - \varepsilon, \infty)$, we want $\#\{\lambda_i > 2 - \varepsilon\} = \mathcal{O}(1)$

► We choose $\varepsilon = cN^{-2/3}$ so that $N\varepsilon^{3/2} = \mathcal{O}(1)$ and

$$\#\{\lambda_i > 2 - cN^{-2/3}\} = \#\{N^{2/3}(\lambda_i - 2) > c\} = \mathcal{O}(1)$$

► This suggests to study the fluctuations of

$$N^{2/3}(\lambda_{\max} - 2)$$

► The $N^{2/3}$ normalization is strongly associated to the \sqrt{x} -behaviour of the density at the corresponding edge

Marčenko-Pastur's theorem

Theorem

- ▶ Consider a $N \times n$ matrix \mathbf{X}_N with i.i.d. entries

$$\mathbb{E}X_{ij} = 0, \quad \mathbb{E}|X_{ij}|^2 = \sigma^2.$$

with N and n of the same order and L_N the spectral measure of $\frac{1}{n}\mathbf{X}_N\mathbf{X}_N^*$:

$$c_n \triangleq \frac{N}{n} \xrightarrow{n \rightarrow \infty} c \in (0, \infty), \quad L_N = \frac{1}{N} \sum_{i=1}^N \delta_{\lambda_i(n^{-1}\mathbf{X}_N\mathbf{X}_N^*)}.$$

- ▶ Then **almost surely** (= for almost every realization)

$$L_N \xrightarrow{N, n \rightarrow \infty} \mathbb{P}_{\tilde{\text{MP}}} \quad \text{in distribution}$$

where $\mathbb{P}_{\tilde{\text{MP}}}$ is **Marčenko-Pastur** distribution:

$$\mathbb{P}_{\tilde{\text{MP}}}(dx) = \left(1 - \frac{1}{c}\right)_+ \delta_0(dx) + \frac{\sqrt{[(\lambda^+ - x)(x - \lambda^-)]_+}}{2\pi\sigma^2xc} dx$$

$$\text{with } \begin{cases} \lambda^- & = \sigma^2(1 - \sqrt{c})^2 \\ \lambda^+ & = \sigma^2(1 + \sqrt{c})^2 \end{cases}$$

Simulations vs $\hat{M}P$ distribution

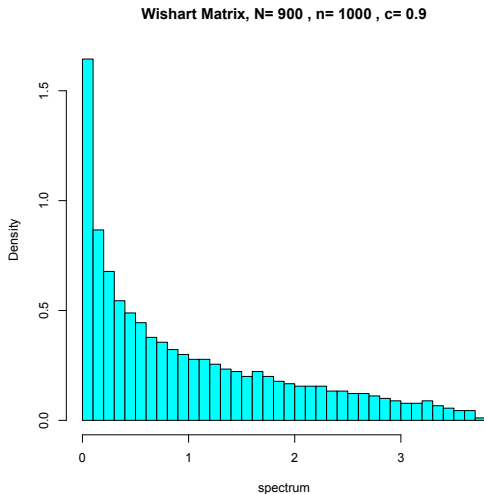


Figure: Histogram of $\frac{1}{n} \mathbf{X}_N \mathbf{X}_N^*$, $\sigma^2 = 1$

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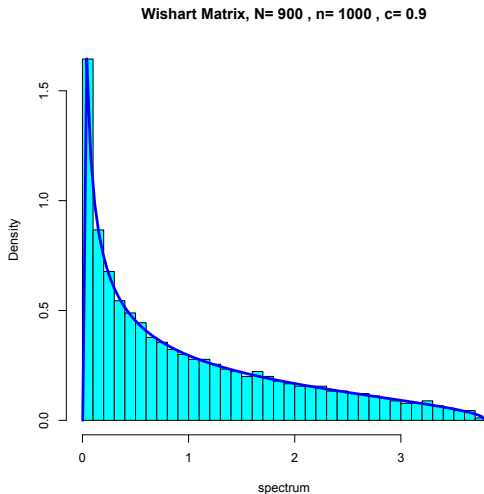


Figure: Marčenko-Pastur distribution for $c = 0.9$

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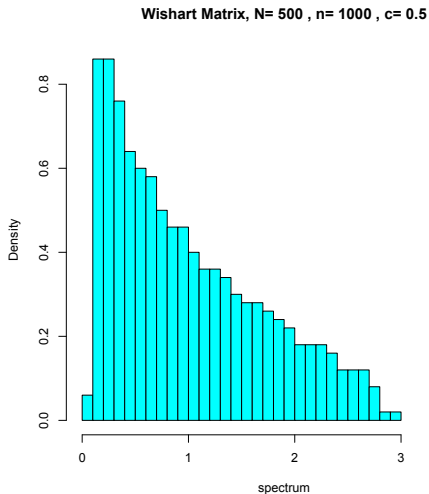


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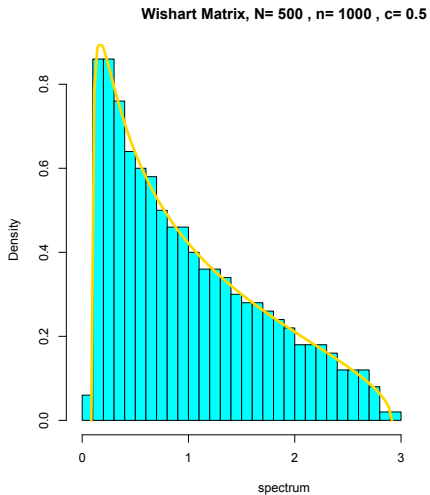


Figure: Marčenko-Pastur distribution for $c = 0.5$

Simulations vs $\hat{M}P$ distribution

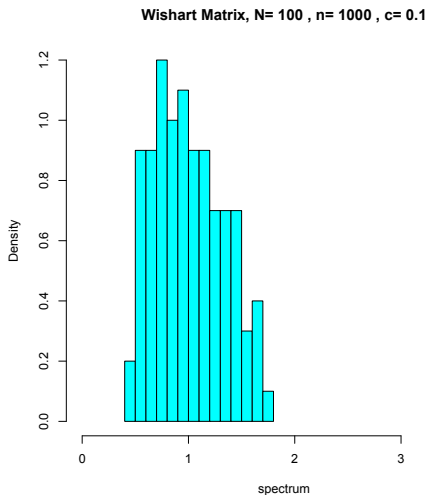


Figure: Histogram of $\frac{1}{n} \mathbf{X}_N \mathbf{X}_N^*$, $\sigma^2 = 1$

Simulations vs $\hat{M}P$ distribution

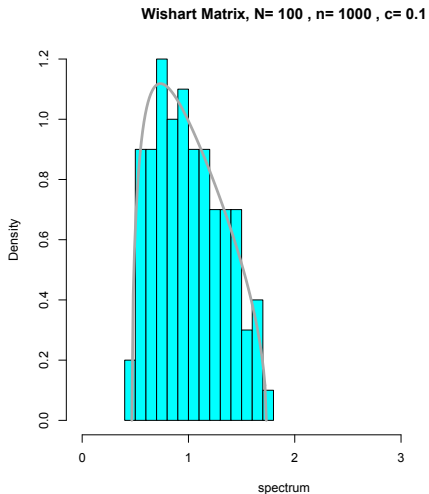


Figure: Marčenko-Pastur distribution for $c = 0.1$

Simulations vs $\hat{M}P$ distribution

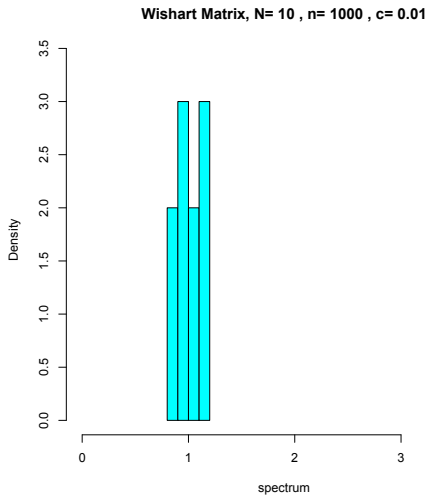


Figure: Histogram of $\frac{1}{n} \mathbf{X}_N \mathbf{X}_N^*$, $\sigma^2 = 1$

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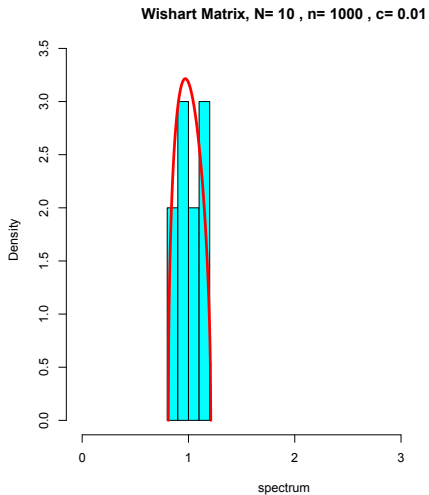


Figure: Marčenko-Pastur distribution for $c = 0.01$

Theorem (fluctuations of λ_{\max} and Tracy-Widom distribution)

$$\frac{N^{2/3}}{\Theta_N} \left\{ \lambda_{\max} \left(\frac{1}{n} \mathbf{X}_N \mathbf{X}_N^* \right) - \sigma^2 (1 + \sqrt{c_n})^2 \right\} \xrightarrow[N, n \rightarrow \infty]{\mathcal{L}} \mathbb{P}_{\text{TW}}$$

where

$$c_n = \frac{N}{n} \quad \text{and} \quad \Theta_N = \sigma^2 (1 + \sqrt{c_n}) \left(\frac{1}{\sqrt{c_n}} + 1 \right)^{1/3}$$

Remark

In this case, we also have Gaussian fluctuations of the linear statistics in the same scale as for the other models.

Non-Hermitian Matrices and the circular law

Matrix model

Let \mathbf{X}_N be a $N \times N$ matrix

$$\mathbf{X}_N = \begin{pmatrix} X_{11} & \cdots & X_{1N} \\ \vdots & & \vdots \\ X_{N1} & \cdots & X_{NN} \end{pmatrix}$$

- ▶ Consider matrix $\mathbf{Y}_N = \frac{1}{\sqrt{N}} \mathbf{X}_N$
- ▶ Beware that the eigenvalues are **complex!**

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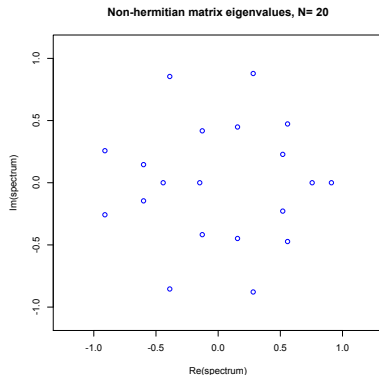


Figure: Distribution of \mathbf{Y}_N 's eigenvalues

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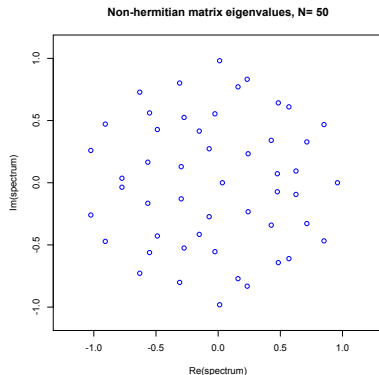


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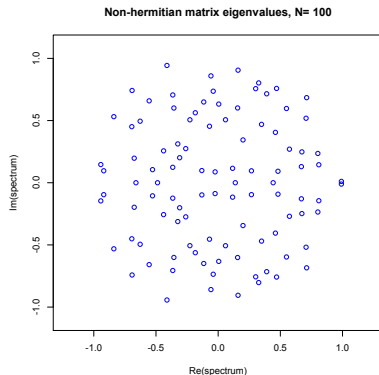


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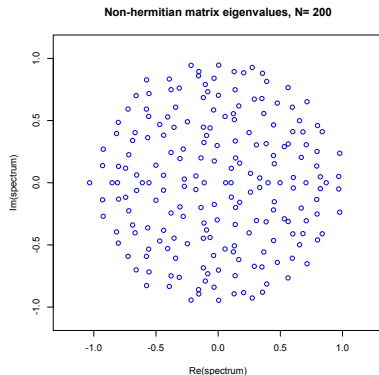


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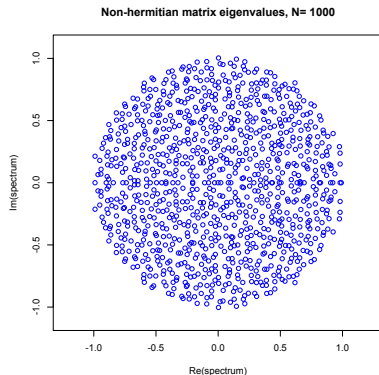


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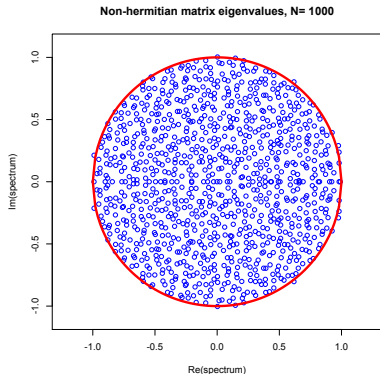


Figure: The circular law (in red)

Theorem: The Circular Law (Ginibre, Mehta, Girko, Tao & Vu, etc.)

The spectrum of \mathbf{Y}_N converges to **the uniform probability on the disc**

The circular law

Theorem

Assume that $\mathbb{E}(X_{ij}) = 0$ and $\mathbb{E}|X_{ij}|^2 = \sigma^2$. Let $Y_N = \frac{X_N}{\sqrt{N}}$, then

$$\text{a.s.} \quad \mu_N^Y = \frac{1}{N} \sum_{i=1}^N \delta_{\lambda_i(Y_N)} \xrightarrow[N \rightarrow \infty]{\text{weakly}} \frac{1}{\pi \sigma^2} \mathbf{1}_{\{x^2 + y^2 \leq \sigma^2\}} dx dy$$

The spectral radius

Assume $\mathbb{E}|X_{ij}|^4 < \infty$ and consider the spectral radius

$$\rho(Y_N) = \max\{|\lambda|, \lambda \in \text{spectrum}(Y_N)\}$$

Then

$$\rho(Y_N) \xrightarrow[N \rightarrow \infty]{\text{a.s.}} \sigma$$

Link for May's criteria for stability

$\mathcal{J} = -I_n + \mathbf{X}_N$, with X_{ij} is $\text{Ber}(C) \cdot \mathcal{N}(0, \sigma^2)$, so that the equilibrium is stable iff

$$\sqrt{NC}\sigma < 1.$$

Fluctuations in the non-Hermitian case

Fluctuations of linear statistics (Rider, Silverstein)

Under appropriate assumptions on the moments of X_{12} , for any analytic function f ,

$$N \left(\frac{1}{N} \sum_{i=1}^N f(\lambda_i(\mathbf{Y}_N)) - f(0) \right)$$

converges to a Gaussian variable, with mean zero and variance $\frac{1}{\pi} \int |f'(z)|^2 dz$.

Fluctuations of spectral radius (Kostlan)

If X_{11} is a Gaussian random variable (Ginibre matrix), then $\rho(\mathbf{Y}_N)$ has Gumbel fluctuations. More precisely, if $c_N := \log N - 2 \log \log N - \log(2\pi)$, for any $x \in \mathbb{R}$,

$$\lim_{N \rightarrow \infty} \mathbb{P} \left(2\sqrt{Nc_N} \left(\rho(\mathbf{Y}_N) - \left(1 + \frac{1}{2} \sqrt{\frac{c_N}{N}} \right) \right) \leq x \right) = e^{-e^{-x}}.$$

Another model of non-Hermitian matrices : elliptic law

Model

- ▶ the diagonal entries are iid and independent of the rest of the matrix, $EX_{11} = 0$ and $EX_{11}^2 < \infty$,
- ▶ for $i \neq j$, the pair (X_{ij}, X_{ji}) is independent of the rest of the matrix, iid

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- ▶ for $i \neq j$, the pair (X_{ij}, X_{ji}) is independent of the rest of the matrix, iid
- ▶ $EX_{12} = EX_{21} = 0$, $EX_{12}^2 = EX_{21}^2 = 1$, $E|X_{12}|^4 < \infty$, $E|X_{21}|^4 < \infty$, $E(X_{12}X_{21}) = \rho \in (-1, 1)$.

As before, we set $\mathbf{Y}_N = \frac{1}{\sqrt{N}} \mathbf{X}_N$

Global behavior (Fyodorov, Girko, Khoruzhenko, Naumov, etc.)

The empirical distribution of the eigenvalues converges to the uniform distribution on the ellipse :

$$\mathcal{E}_\rho := \left\{ (x, y) \in \mathbb{R}^2; \frac{x^2}{(1+\rho)^2} + \frac{y^2}{(1-\rho)^2} \leq 1 \right\}$$

Local behavior (Renfrew)

For any $\delta > 0$, a.s. for N large enough, there is no eigenvalue of \mathbf{Y}_N outside a δ -neighborhood of \mathcal{E}_ρ .

Spiked models

General idea

If you take your favorite random matrix model and add a deterministic matrix with low rank (that is the rank stays bounded as N goes to infinity), then

- ▶ the global behavior is not affected
- ▶ the behavior of extreme eigenvalues can be affected drastically if the perturbation is strong enough.

There is a huge literature on the subject...

The simplest case can be stated as follows :

Wigner case

We denote by Θ_N a rank one matrix with eigenvalue θ and delocalized eigenvectors. For example,

$$\Theta_N := \theta \mathbf{u} \mathbf{u}^*,$$

with u uniform on the sphere in dimension N . Let \mathbf{Y}_N be a Wigner matrix as defined earlier and $\lambda_{\max} := \lambda_{\max}(\mathbf{Y}_N + \Theta_N)$.

Then,

- ▶ if $\theta \leq 1$, $N^{2/3} (\lambda_{\max} - 2)$ converges to a Tracy-Widom distribution,
- ▶ if $\theta > 1$, $\sqrt{N} (\lambda_{\max} - (\theta + \frac{1}{\theta}))$ converges to a Gaussian distribution.

Simulations

Simulations

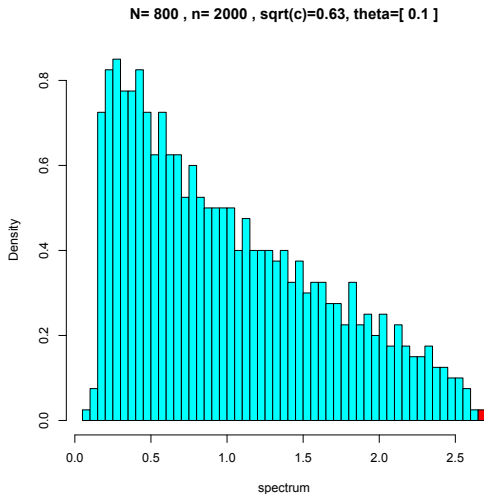


Figure: Spiked model - strength of the perturbation $\theta = 0.1$

Simulations

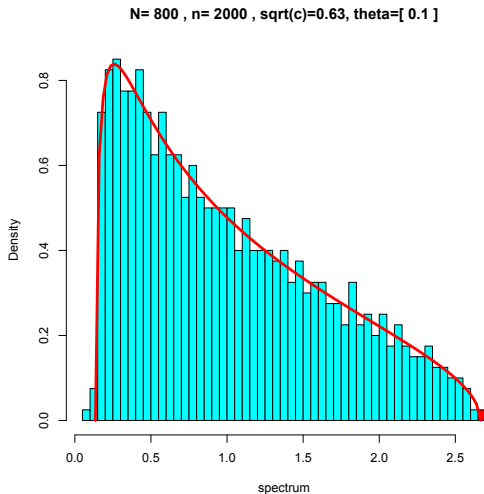


Figure: Spiked model - strength of the perturbation $\theta = 0.1$

Simulations

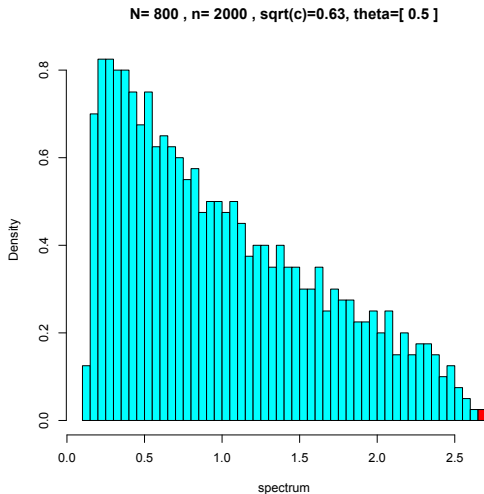


Figure: Spiked model - strength of the perturbation $\theta = 0.5$

Simulations

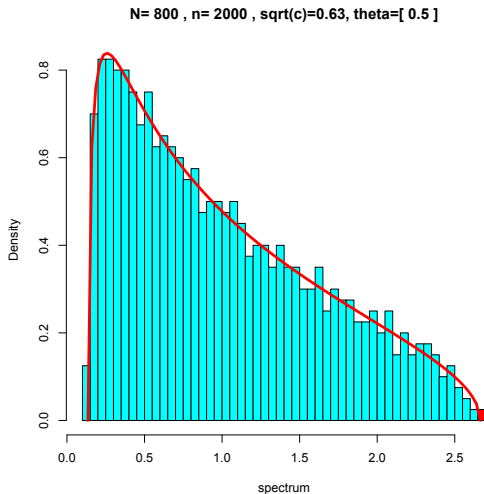


Figure: Spiked model - strength of the perturbation $\theta = 0.5$

Simulations

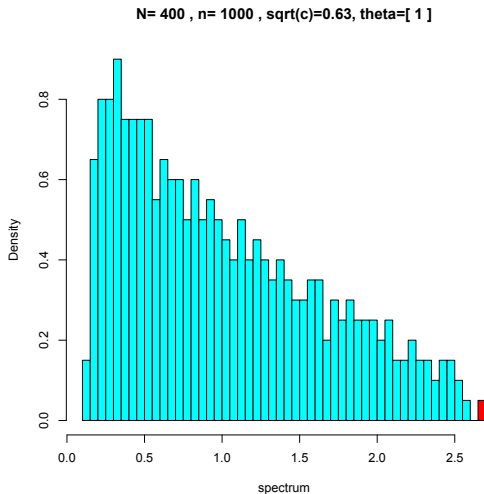


Figure: Spiked model - strength of the perturbation $\theta = 1$

Simulations

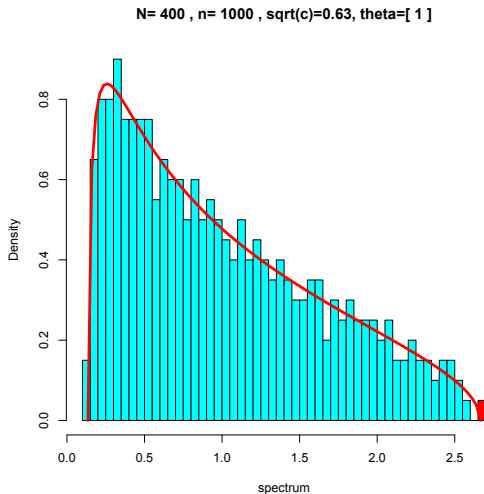


Figure: Spiked model - strength of the perturbation $\theta = 1$

Simulations

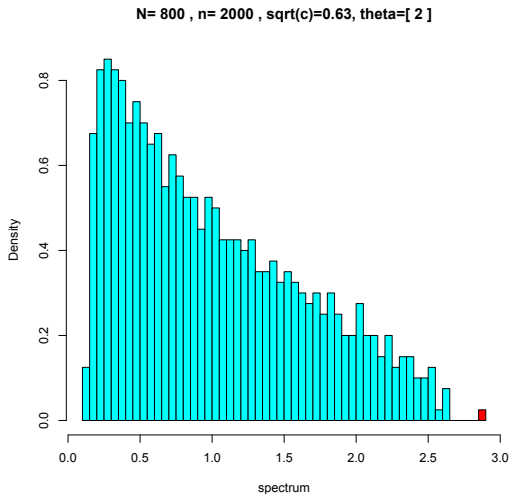


Figure: Spiked model - strength of the perturbation $\theta = 2$

Simulations

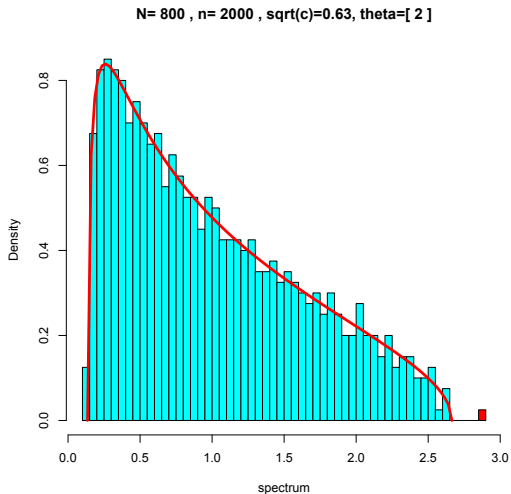


Figure: Spiked model - strength of the perturbation $\theta = 2$

Simulations

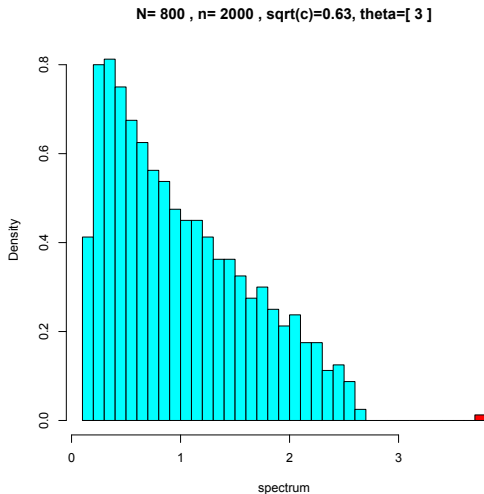


Figure: Spiked model - strength of the perturbation $\theta = 3$

Simulations

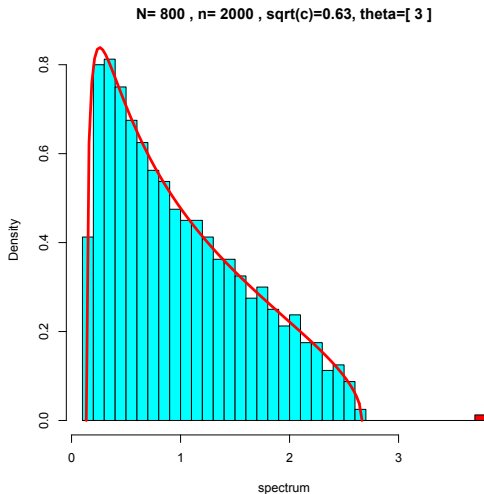


Figure: Spiked model - strength of the perturbation $\theta = 3$

To conclude

Many interesting questions in RMT arise from ecological models.

1. Normalization and feasibility,
2. General interaction matrix models

Network type	Statistical Features	RMT Results
random	Z_{ij} i.i.d. and $\mathbb{E}Z_{ij} = 0$	Circular law
structured	$Z_{ij} = 0$ for $(i, j) \in \mathcal{S}$	Sparse variance profiles / open
mutualistic	$\mathbb{E}Z_{ij} > 0$	open
predator-pray	$Z_{ij} = -Z_{ji}$	open
density-dependent	$\mathcal{J} = DZ$	open

Figure: Various types of ecological networks

Thank you for your attention !